

# Algorithms for lattice QCD I

Stefan Schaefer

NIC, DESY

Kolkata Lattice Gauge Theory School

## Definition

Wilson '74

Using a discrete space-time QCD can be defined.

## Computational method

Path integral  $\rightarrow$  high dimensional integral

Computation of integrals by Monte Carlo.

E.g. hadron masses, decay constants, ...

## Algorithms

Many calculations computationally very expensive.

Massively parallel computers employed.

Need to work on simulation setup, algorithms, ...

Many choices.

Physics understanding  $\leftrightarrow$  Numerical setup.

## Field update

Integral over field variables  $\rightarrow$  **sum** over field configurations

- How to efficiently generate these fields.

- Transformation of action into form amenable to numerical treatment.

- Numerical solution of diff. eq.

## Solution of Dirac equation

Needed in field update and fermionic observables.

Out-of-the-box algorithms, e.g., the conjugate gradient, perform badly

$\rightarrow$  Need to take physics into account

## **Monday**

Introduction

Markov Chain Monte Carlo

The HMC algorithm

## **Tuesday**

MD integrators

Fermions in QCD simulations

Mass preconditioning

## **Wednesday**

Other fermion methods (RHMC, DD-HMC)

Solution of the Dirac equation

## **Thursday**

Solution of the Dirac equation II

Local deflation

Methods to compute hadron observables

## **Friday**

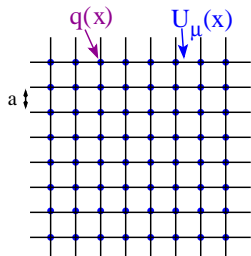
Autocorrelations

General reading:

M. Lüscher

Computational strategies in lattice QCD

Les Houches 2009



## Computation of path integral

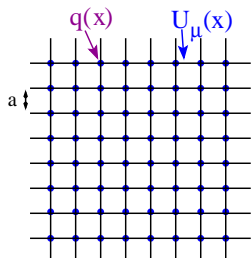
$$\langle A \rangle = \frac{1}{Z} \int \prod_{x,\mu} dU_{x,\mu} e^{-S[U]} A[U]$$

with

$$Z = \int \prod_{x,\mu} dU_{x,\mu} e^{-S[U]}$$

One SU(3) integration variable for each link.

# Goal



$$\langle A \rangle = \frac{1}{Z} \int \prod_{x,\mu} dU_{x,\mu} e^{-S[U]} A[U]$$

One SU(3) integration variable for each link.

$128 \times 64^3$  lattice  $\rightarrow 1.3 \cdot 10^8$  links

Classical numerical quadrature would need  
 $N^{\text{\#variables}}$  function evaluations

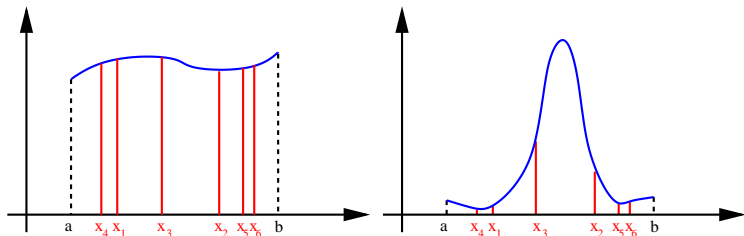
# Monte Carlo

General idea of Monte-Carlo integration

$$\frac{1}{b-a} \int_a^b dx f(x) \approx \frac{1}{N} \sum_{i=1}^N f(x_i)$$

with randomly chosen points  $x_i$  in the integration region

- good idea, if  $f(x)$  approximately constant  
⇒ small fluctuations in  $f(x_i)$ .
- For a given realization of the  $N$  points  $x_i$ , this is an **unbiased estimator** of the integral





# Error of a MC simulation

$$\tilde{F}_j = \frac{1}{N} \sum_{i=1}^N f(x_i^j)$$

Index  $j$  labels the repetition of the “experiment”.

Unbiased = gives correct result *on average*

$$F = \langle\langle \tilde{F} \rangle\rangle$$

$\langle\langle \cdot \rangle\rangle$  average over realizations of the  $x_i$

Typical deviation  $\rightarrow$  variance of *this estimator*.

$$\begin{aligned} \langle\langle (\tilde{F} - F)^2 \rangle\rangle &= \frac{1}{N^2} \sum_{ik} \langle\langle F_i F_k \rangle\rangle - F^2 \\ &= \frac{1}{N} \left( \sum_i \langle\langle (F_i)^2 \rangle\rangle - F^2 \right) = \text{var}(F)/N \end{aligned}$$

Error decreases as  $1/\sqrt{N}$

# Note of caution

This theoretical analysis assumes knowledge of two quantities

$$\bar{F} = \int dx f(x) \quad \text{and} \quad \text{var}(f) = \int dx (f(x) - \bar{f})^2$$

What you get from the Monte Carlo are **estimators** of these quantities.

The analysis is correct for  $N \rightarrow \infty$ ;  $\infty$  is a large number.

These estimators might have significant errors, which are hard to get from the MC. You might also just have been unlucky.

**To a certain extent, practical Monte Carlo is an art and requires careful inspection of the results.**

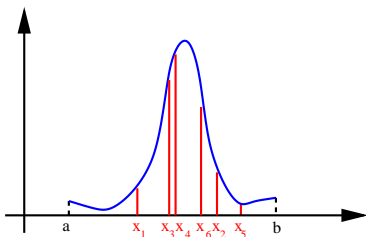
# Importance sampling

Estimator correct up to  $\sqrt{\text{var}(f)/N} \rightarrow$  reduce variance.

$$\int \mathbf{d}x f(x) = \int \rho(x) \mathbf{d}x \frac{f(x)}{\rho(x)} = \left[ \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{\rho(x_i)} \right] (1 + \mathcal{O}(N^{-1/2}))$$

with points  $x_i$  chosen according to  $\rho$ .

- Choose points according to probability distribution similar to function to be integrated
- optimal, if distribution  $\propto |f(x)| \dots$  need  $\int dx |f(x)|$



# Markov Chain Monte Carlo

Problem of “straight” Monte Carlo is to find a *normalized* probability density  $P(x)$

## Solution

Use a method which only needs **relative** probabilities

Construct a sequence of points

$$x_1 \rightarrow x_2 \rightarrow x_3 \rightarrow \cdots \rightarrow x_N$$

using a transition probability  $T(x_{i+1} \leftarrow x_i)$

Analysis using arguments of a  $N$  sets of such chains.

# Properties of $T$

For any given pair of points  $x_1$  and  $x_2$

$$T(x_2 \leftarrow x_1)$$

with the following properties

(A) **Stability**

$$P(x') = \int dx P(x) T(x' \leftarrow x)$$

(B) **Normalization**

$$\int dx' T(x' \leftarrow x) = 1$$

(C) **Ergodicity**

$$T(x' \leftarrow x) > 0 \text{ for each pair } x, x'$$

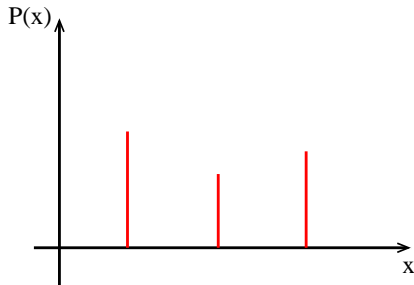
**No reference to absolute normalization of  $P$**

# Analysis of MCMC

**For the sake of simplicity, consider discrete state space.**

Integrals  $\rightarrow$  sums.

Example for a single variable with three possible values:



$T(x' \leftarrow x)$  is a matrix acting in the space of states.

For the example a  $3 \times 3$  matrix, because  $x$  can take 3 values.

A probability distribution is a normalized vector in this space.

Imagine having an ensemble of points  $x_i$  distributed according to  $P(x)$ .

Condition (A) reads

$$P = TP$$

→  $P$  is eigenvector of  $T$  with eigenvalue 1.

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## Theorem of Frobenius–Perron

For a matrix with the properties A–C the following holds

- There is exactly one eigenvalue  $\lambda = 1$ .
- All eigenvalues  $\lambda$  have  $|\lambda| \leq 1$ .

# Convergence of Markov Process

Given any starting distribution  $P_0$  repeated application of  $T$  leads to exponential convergence to desired distribution

$$\begin{aligned} T^n P_0 &= \sum_{i=0}^N \lambda_i^n (\Psi_i, P_0) \Psi_i \\ &= \sum_{i=0}^N e^{\log |\lambda_i| n} (\Psi_i, P_0) \frac{\lambda_i}{|\lambda_i|} \Psi_i \\ &\propto P + O(e^{-n/\tau_1}) \end{aligned}$$

$\tau_i = 1/\log |\lambda_i|$  are the exponential autocorrelation times

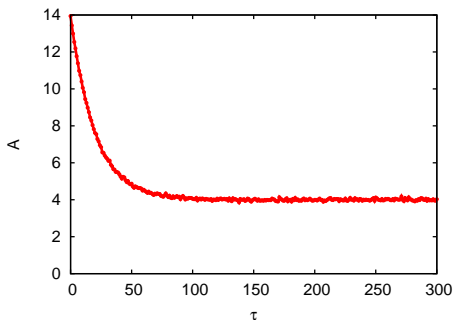
A nicer interpretation in terms of single exponentials can be given with detailed balance, see later.



# Practical MCMC

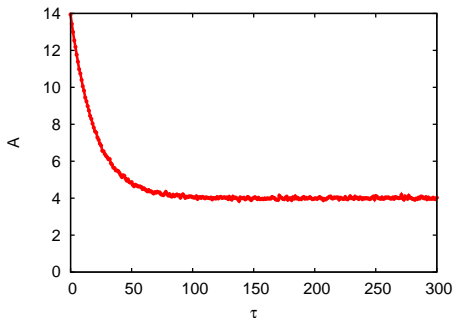
Start with one (or a few) points  
Averages of Monte Carlo time

$$\langle f \rangle = \frac{1}{N} \sum_{i=1}^N f(x_i)$$



Cut away first *thermalization* phase.

# Time evolution



$$\langle f \rangle = \frac{1}{N} \sum_{t=1}^N f(x_t) (1 + \mathcal{O}(1/\sqrt{N}))$$

At the beginning large contributions from  $\Psi_{i>0}$ .

Formally, they are suppressed by  $1/N$ , but they are best not considered.

A valid simulation must have  $n \gg 1/\tau_1$

# Autocorrelations

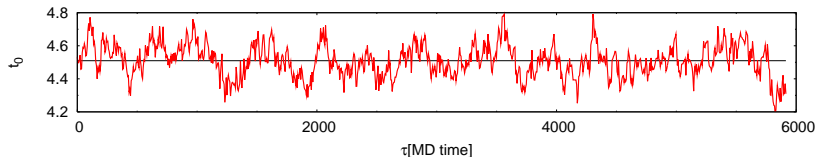
There is a price to pay:

Subsequent points  $x_i$ /configurations are in general not independent.

$$\langle\langle f(x_i)f(x_{i+1}) \rangle\rangle \neq 0$$

This needs to be taken into account in the *analysis* of the data.

It also can make simulations exceedingly expensive.



# Construction of $T$

Constructing a valid  $T$  seems a daunting task.

## Metropolis–Hastings method

Metropolis et al'53

- Symmetric proposal (trans. prob. w/  $P_0(x) = \text{const}$ )

$$W(x' \leftarrow x) = W(x \leftarrow x')$$

- Acceptance step

$$P_{acc}(x', x) = \min\left[1, \frac{P(x')}{P(x)}\right]$$

Next point is  $x'$  with propability  $P_{acc}$ , else  $x$

Total transition propability

$$T(x' \leftarrow x) = W(x' \leftarrow x)P_{acc}(x', x) + \delta_{xx'} \sum_{\tilde{x}} (1 - P_{acc}(\tilde{x}, x))W(\tilde{x} \leftarrow x)$$

$$T(x' \leftarrow x) = W(x' \leftarrow x)P_{acc}(x', x) + \delta_{xx'} \left[ \sum_{\tilde{x}} (1 - P_{acc}(\tilde{x}, x)) W(\tilde{x} \leftarrow x) \right]$$

**(A) Stability**

$$\begin{aligned} & \sum_x P(x) T(x' \leftarrow x) \\ &= \sum_x W(x' \leftarrow x) P(x) P_{acc}(x', x) + P(x') \sum_{\tilde{x}} (1 - P_{acc}(\tilde{x}, x)) W(\tilde{x} \leftarrow x) \\ &= \sum_x W(x \leftarrow x') P(x') P_{acc}(x, x') + P(x') (1 - \sum_{\tilde{x}} P_{acc}(\tilde{x}, x) W(\tilde{x} \leftarrow x)) = P(x') \end{aligned}$$

**(B) Normalization**

$$\sum_{x'} T(x' \leftarrow x) = 1$$

# Practical implementation

We want to generate the sequence

$$x_1 \rightarrow x_2 \rightarrow x_3 \rightarrow \dots$$

and arrived at  $x_j$

- Make a proposal  $y$  according to probability  $W(y \leftarrow x_j)$
- Compute  $P_{\text{acc}}$  and draw random number  $0 \leq r < 1$

$$x_{j+1} = \begin{cases} y & \text{if } r < P_{\text{acc}} \\ x_j & \text{else} \end{cases}$$

Good proposal

- Easy to generate
- High acceptance probability

# Generic issues of the method

The step size can be rather limited

## Small steps

$P(x') \approx P(x) \rightarrow P_{\text{acc}}$  high

But it takes many steps to sample the whole integration space

Large autocorrelations

## Large steps

Small autocorrelations

Feasible only if an update leading to reasonable  $P_{\text{acc}}$  can be found

Can be difficult to compute proposal.

In high dimensional spaces, it can be very difficult to argue, what “large step” and “small step” even mean.

# Summary of part I

Monte Carlo is a method to numerically compute high-dimensional integrals.

Integral  $\rightarrow$  average over sample points.

Sample points need to be chosen in region of high probability.

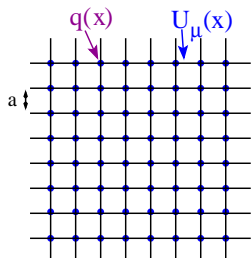
Problem of constructing a normalized distribution

$\rightarrow$  Markov Chain Monte Carlo

No normalized distribution needed.

Deal with autocorrelations.





One point  $x \rightarrow$  one value for each link variable  
 $\rightarrow$  one field configuration

Instead of  $x$  we will therefore now use  $U$  which are in  $SU(3)$ .

$$\langle \mathbf{A} \rangle = \frac{1}{Z} \int \prod_{x,\mu} dU_{x,\mu} e^{-S[U]} \mathbf{A}[U]$$

with

$$Z = \int \prod_{x,\mu} dU_{x,\mu} e^{-S[U]}$$

Normalized probability density

$$P[U] = \frac{1}{Z} e^{-S[U]}$$

Ratios in probabilities

→

Need to evaluate *differences* in the action  $S[U]$

**While this choice of  $P[U]$  seems natural, it is not unique**

Basically two types of algorithms

### Single link updates

Of the  $4V$  links, only one is changed at a time.

In each step  $S[U] - S[U']$

Possible if this is an  $O(1)$  operation.

With dynamical fermions, estimating change in action is a global  $O(V)$  operation

Method of choice in pure gauge theory.

No systematic study.

Basically two types of algorithms

## Updates based on Molecular Dynamics

Based on ideas from classical mechanics.



Field configuration position

Introduce momenta  $\rightarrow$  equations of motion.

Updates keep probability constant (micro-canonical)

Solves problem of finding a good proposal in  
Metropolis-Hastings procedure

## Extended field space

$$Z = \int [dU][d\pi] e^{-\frac{1}{2}(\pi, \pi) - S[U]}$$

Expectation values of observables  $A[U]$  remain the same.

Momenta  $\pi = \pi^a T^a \in \mathfrak{su}(N)$ ,  $\pi^a \in \mathbb{R}$

$$(\pi, \pi) = \sum_{x, \mu} \pi_{x, \mu}^a \pi_{x, \mu}^a$$

## Updates

Make updates in this extended phase space.  
 $\Rightarrow$  updates for  $U$  fields.

Essential update step for the gauge fields:  $(\pi, U) \rightarrow (\pi', U')$

## Hamilton's equations of motion

Hamiltonian

$$H[\pi, U] = \frac{1}{2}(\pi, \pi) + S[U]$$

E.o.m.

$$\dot{U}_{x,\mu} = \pi_{x,\mu} U_{x,\mu}$$

$$\dot{\pi}_{x,\mu} = -F_{x,\mu} ,$$

$$F_{x,\mu}^a = \frac{\partial S(e^\omega U)}{\partial \omega^a(x, \mu)}$$

By **Liouville's theorem**, the classical dynamics

$$(\pi, U) \rightarrow (\pi', U')$$

maps areas of equal likelihood into each other.

*Energy conservation*

$$\frac{d}{d\tau} H = 0$$

Boltzmann factor  $e^{-H}$  is constant.

*Conservation of phase space*

**An exact solution of the E.o.m is a valid update.**

Fundamentally different from Metropolis–Hastings

In practice: integration errors

## Momentum Heatbath

Refresh momenta  $\pi$  (Gaussian random numbers)

## Molecular Dynamics

Solve numerically MD equations for some MC time  $\tau$  (trajectory) deriving from Hamiltonian  $H = \frac{1}{2}(\pi, \pi) + S[U]$ .



## Acceptance Step

Correcting for inaccuracies in integration.



# Metropolis

Different viewpoint:

The molecular dynamics as the symmetric proposal in Metropolis.

This solves the problem of the inexact integration.

Need a **symplectic integrator**, i.e. area conserving and reversible.

Reversibility haunts computer implementations.  
No good theory for this.

## Acceptance step

Molecular dynamics  $(\pi, U) \rightarrow (\bar{\pi}, \bar{U})$

$$P_{\text{acc}} = \min(1, e^{-(H(\bar{\pi}, \bar{U}) - H(\pi, U))})$$

Exact solution of MD equations has  $\Delta H = 0$

$\Rightarrow$  always accepted

$\Delta H$  needs to be  $O(1)$  for good acceptance.  
Difficult to achieve on large volume.

$$Z = \int [dU][d\pi] e^{-\frac{1}{2}(\pi, \pi) - S[U]}$$

## Momenta: Heatbath

- $(\pi, \pi) = \sum_{x, \mu} |\pi_{x, \mu}^a|^2$   
 $\pi^a$  are Gaussian random numbers.
- Normalization is known.

Just for  $\pi$  this is the optimal update

no correlation to previous config.

Also true for combined system? (Kramers' rule, ...)

**Momentum heat-bath is the only source of randomness.**

Makes algorithm ergodic.

Problems with ergodicity from  $S = \infty$  surfaces.

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The original hope was that a trajectory constitutes a macroscopic update.

Free field theory

$$\ddot{U} = \dot{\pi} = \frac{\delta S}{\delta U}$$

→ Monte Carlo Time  $\tau \propto 1/a$

In the classic algorithm trajectory length is scaled with  $1/a$ .

In number of updates, autocorrelations should stay the same.

M.Lüscher, S.S.'11

This does not apply in interacting theory  
→ Wednesday

Virota,Sommer, S.S.'10

Still, rather long trajectories are a good idea,  $\tau \sim 2$ .

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For exercises in  $\phi^4$  theory: S.S., Les Houches 2009,  
available at [NIC@DESY](mailto:NIC@DESY) website.